

Determining Noise Levels in Blurry Image Data

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Suppose we have data of the form

$$b = h * u + \epsilon, \quad (1)$$

where u is an image we want to recover, h is a point spread function (PSF), and ϵ is a noise term. Suppose $\epsilon \sim N(0, \sigma^2 I)$, where σ is unknown. Let \mathcal{F} denote the unitary discrete Fourier transform operator. Let $\hat{x} = \mathcal{F}x$ for an image x . Then

$$\hat{b} = \hat{h} \cdot \hat{u} + \hat{\epsilon}. \quad (2)$$

Proposition 1. *Suppose $\epsilon \sim N(0, \sigma^2 I)$. Then*

$$\hat{\epsilon} = |\hat{\epsilon}| e^{i\theta}, \quad (3)$$

where $|\hat{\epsilon}| \sim N(0, \sigma^2 I)$ and $\theta \sim U([- \pi, \pi])$.

This result is explained in the abstract of this article [1].

Proof. Before going into the rigorous proof, we provide some leading results. First

$$\begin{aligned} \mathbb{E} \hat{\epsilon}_k &= N^{-1/2} \mathbb{E} \sum_{j=0}^{N-1} \epsilon_j e^{-i2\pi k j / N} \\ &= N^{-1/2} \sum_{j=0}^{N-1} \mathbb{E} \epsilon_j e^{-i2\pi k j / N} = 0. \end{aligned} \quad (4)$$

Next

$$\begin{aligned} \mathbb{E} |\hat{\epsilon}_k|^2 &= N^{-1} \mathbb{E} \sum_{m=0}^{N-1} \sum_{n=0}^{N-1} \epsilon_m \epsilon_n e^{-i2\pi k(m-n)/N} \\ &= N^{-1} \sum_{m=0}^{N-1} \mathbb{E} \epsilon_m^2 = \sigma^2. \end{aligned} \quad (5)$$

So the mean and variance match the claims of the proof.

In general, for a Gaussian variable X with mean μ covariance matrix σ , the characteristic function is given by

$$\phi_X(t) = \exp\left(i\mu t - \frac{t^2 \sigma^2}{2}\right),$$

and for any constants a, b and independent random variable X, Y the characteristic function for $aX + bY$ is

$$\phi_{aX+bY}(t) = \phi_X(at) \phi_Y(bt).$$

Therefore, combining these two facts, the characteristic function for $X_k = \text{Re}\{\hat{\epsilon}_k\}$ is

$$\phi_{X_k}(t) = \exp\left(-\frac{t^2\sigma^2}{2N} \sum_{j=0}^{N-1} \cos^2(2\pi kj/N)\right) \quad (6)$$

Write \cos^2 as

$$\cos^2(2\pi kj/N) = \left(\frac{e^{i2\pi kj/N} + e^{-i2\pi kj/N}}{2}\right)^2 = \left(\frac{e^{i4\pi kj/N} + e^{-i4\pi kj/N} + 2}{4}\right).$$

Substituting this into the sum in (6) obtains

$$\sum_{j=0}^{N-1} \cos^2(2\pi kj/N) = N/2,$$

therefore the characteristic for X_k is

$$\exp\left(-\frac{t^2\sigma^2}{4}\right),$$

which is the characteristic of a mean zero normal distribution with variance $\sigma^2/2$. Repeating this result on the imaginary part ($Y_k = \text{Im}\{\epsilon_k\}$) of the Fourier coefficients with sines almost completes the proof. What we have shown is that

$$X_k \sim N(0, \sigma^2/2) \quad \text{and} \quad Y_k \sim N(0, \sigma^2/2).$$

The remainder of the proof would be to show then that

$$\hat{\epsilon}_k = X_k + iY_k$$

satisfies the statement of the proposition. The claim about the squared magnitude is straightforward. The claim about the phase is not clear to me how to show. \square

Now given b , we can estimate σ^2 in the following way. The Fourier transform \hat{b} is composed of two parts $\hat{\epsilon}$ and $\hat{h} \cdot \hat{u}$. The first term is described in the previous proposition. The second term is something which has been low pass filtered by h . Therefore, what remained at the high pass regions should be dominated by $\hat{\epsilon}$. Therefore, we take \hat{b} , isolate the high wave numbers to be some set say S , and take the average over the squared terms in S to estimate σ^2 :

$$\hat{\sigma}^2 = |S|^{-1} \sum_{k \in S} |\hat{b}_k|^2.$$

Be careful not to estimate σ by just averaging the magnitudes. If one wanted to do that, then it should be done with the following formula:

$$\hat{\sigma} = |S|^{-1} \sqrt{\frac{\pi}{2}} \sum_{k \in S} |\hat{b}_k|.$$

The scaling factor comes from the fact that for a random variable $X \sim N(0, \sigma^2)$, it can be shown that

$$\mathbb{E}|X| = \sqrt{\frac{2}{\pi}}\sigma.$$

Median Absolute Deviation

A more robust σ estimation is given by the median absolute deviation (MAD):

$$\hat{\sigma} = \frac{1}{0.6745} \text{median}_{k \in S} |\hat{b}_k|.$$

The median estimation is more robust to outliers, and for a normally distributed data set $\{x_i\}_i$ with mean 0 and variance 1, 50% of the distribution is on the interval $[-.6745, .6745]$ (i.e. $\phi(.6745) - \phi(-.6745) \approx .5$), hence

$$\mathbb{E}[\text{median}_i |x_i|] = .6745.$$

Fourier transform of Gaussian

This seems like a good time for a formal proof of the well known fact: *the Fourier transform of a Gaussian is a Gaussian*. This is seen in the earlier derivation, where the characteristic function is a Gaussian. I wanted to give the formal proof here:

Proposition 2. *Let*

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{x^2}{2\sigma^2}\right).$$

Then

$$\hat{f}(\xi) = \int_{\mathbb{R}} f(x) e^{-i2\pi x \xi} dx = \exp(-2\sigma^2 \pi^2 \xi^2)$$

Proof. The key to the proof is the completing the square with a complex number:

$$x^2 + i4\pi\sigma^2\xi x = (x + i2\pi\sigma^2\xi)^2 + (2\pi\xi\sigma^2)^2. \tag{7}$$

Using this we obtain

$$\begin{aligned} \hat{f}(\xi) &= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{\mathbb{R}} \exp\left(-\frac{x^2}{2\sigma^2} - i2\pi x \xi\right) dx \\ &= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{\mathbb{R}} \exp\left(-\frac{1}{2\sigma^2} (x^2 + i4\pi\sigma^2 x \xi)\right) dx \\ &= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{\mathbb{R}} \exp\left(-\frac{1}{2\sigma^2} ((x + i2\pi\sigma^2\xi)^2 + (2\pi\xi\sigma^2)^2)\right) dx \\ &= \frac{\exp(-2\sigma^2\pi^2\xi^2)}{\sqrt{2\pi\sigma^2}} \int_{\mathbb{R}} \exp\left(-\frac{1}{2\sigma^2} (x + i2\pi\sigma^2\xi)^2\right) dx. \end{aligned} \tag{8}$$

The remainder of the proof involves showing in the last line that the integral cancels the denominator. Certainly, if $\xi = 0$ this is true. It then suffices to show

$$F(a) = \int_{\mathbb{R}} \exp(-(x + ia)^2) dx$$

is a constant by showing the $F'(a) = 0$.

Note: it is probably much easier to work with $f(x) = e^{-x^2}$ and just change variables later using the scaling properties of the Fourier transform. \square

References

- [1] D. Freedman et al. The empirical distribution of fourier coefficients. *The Annals of Statistics*, 8(6):1244–1251, 1980.